Research Article



Computing Method of Multivariate Process Capability Index Based on Normalized Pretreatment

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Abstract:

For the traditional multi-process capability construction method based on principal component analysis, the process variables are mainly considered, but not the process capability, which leads to the deviation of the contribution rate of principal component. In response to the question, this paper first clarifies the problem from two aspects: theoretical analysis and example proof. Secondly, aiming at the rationality of principal components degree, an evaluation method for pre-processing data before constructing MPCI using PCA is proposed. The pre-processing of data is mainly to standardize the specification interval of quality characteristics making the principal components degree more reasonable and optimizes the process capability evaluation method. Finally, the effectiveness and feasibility of the method are proved by an application example.

Keywords: Multivariate process capability index; The standard range; Contribution degree; Specification intervals

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1. Introduction

Process capability is the consistency index of machining process. In order to measure, evaluate, analyze, and compare process capabilities, quantifiable indicators as Process Capability Index (PCI) are practically used. The concept of process capability ratio are firstly proposed by Juran^[1], who compared process fluctuations with process specification as a quantitative index to evaluate process capability. Other quantifiable indicators that have emerged since then are based on process fluctuations and specifications, such as C_p , C_{pk} and C_{pm} . With the improvement of processing level and increasingly complex parts, the processing parts generally have multiple quality characteristics, and the process capability evaluation needs to consider multiple quality characteristics. Since the concept of a multi-process capability index are proposed by Chan L K et al.^[2], Process Capability Index multivariate analysis model was constructed as the following categories: the multi-process capability index (MPCI) was constructed by region ratio, such as the method for evaluating the process area that are presented by Shahriari et al.^[3], and the spatial MPCI proposed by Wang S. X et al.^[4] to solve the eccentricity of multivariate quality features. MPCI was constructed by a failure rate, such as Wierda et al.^[5] Bothe et al.^[6] and Chen K S et al.^[7]. MPCI was constructed by principal component analysis, such as the process capability index of principal components proposed by Wang KF and Chen JC^[8], Wang FK and Du TC^[9], Ma Yizhong^[10] and Wang Rang CH^[11] to build multi-quality characteristics index. And based on the correlation of the specification intervals, Shinde et al.^[12] and Zhang M et al.^[13] modified the specification interval. MPCI is constructed by vector method, such as a new multivariate capability vector that are proposed by Shahriari and Abdollahzadeh^[14]. In addition to the methods above, there are another method. Pan^[15] proposed two new multi-process capability indices: NMC_p and NMC_{pm}.

In the traditional Principal Component Analysis (PCA) constructing the Multivariate Process Capability Index (MPCI), a few principal components mainly consider process variation but not process capability, which is unreasonable. Therefore, aiming at the rationality of principal components degree, an evaluation method for pre-processing data before constructing MPCI using PCA is proposed.

2. Principal Component Analysis theory

When the mass data of multi-quality feature parts obeys the multivariate normal distribution, the principal component analysis is proposed by Wang KF and Chen $JC^{[8]}$ to construct the multivariate process capability index. They extended the one-unit capability index to pluralism and defined the multivariate process capability indices: MC_p , MC_{pk} , MC_{pm} and MC_{pmk} .

When using the principal component analysis method, a small number of principal components can explain about 90% of the process variation. In terms of construction methods, the method of weighted averaging is used by Ma Yizhong^[10]. Ac-

Copyright © 2019 by author(s) and Viser Technology Pte. Ltd. This is an Open Access article distributed under the terms of the Creative Commons Attribution-NonCommercial 4.0 International License (http://creativecommons.org/licenses/by-nc/4.0/), permitting all non-commercial use, distribution, and reproduction in any medium, provided the original work is properly cited. cording to the fluctuations, for each principal component index, different weights are assigned. Take Cp as an example:

$$MCp = \sum_{i=1}^{\nu} r_i Cp; PC_i$$
(1)

And Wang F K and Du T $C^{[9]}$ and Wang Rang C $H^{[11]}$ use geometric mean, namely:

$$MCp = \left(\prod_{i=1}^{\nu} Cp; PC_i\right)^{1/\nu}$$
(2)

Where

$$Cp; PC_i = \frac{USL_{PC_i} - LSL_{PC_i}}{6\sigma_{PC_i}}$$

Where, $r_i = \lambda_i / tr(\Lambda, tr(\Lambda) = \lambda_1 + \lambda_2 + ... + \lambda_v$, here, $\lambda_1, \lambda_2, ... \lambda_p$ is the feature vector of Σ , $Cp; PC_i$ is the single process capability index of the i-th principal component, and V represents the number of principal components which can explain about 90% of the process variation. The specification range and target value of the principal component are:

$$LSL_{PCi} = u_i'LSL \quad USL_{PCi} = u_i'USL \tag{3}$$

Where, $u_1, u_2, ..., u_p$ is the feature vector of Σ .

Shinde and Khadse^[12] pointed out that the formulas calculated by Wang K F et al.^[8-9] are biased, because in principal component analysis, only the distribution of principal components is independent, the canonical intervals of different principal components are still interrelated, and their viewpoints are verified by examples. At the same time, new specification intervals and multi-process capability indices are defined.

$$V = \{ (y_1, y_2, \dots, y_v) | LSL_X \le Uy \le USL_X, y' = (y_1, y_2, \dots, y_p), y_r = EY_r \} r = v + 1, v + 2, \dots, p$$
(4)

Where, y is the observations of the principal component (PC), V is not a super-rectangular, it is composed of 2p linear inequalities (p is the number of quality features).

$$Mp_{1} = P \begin{cases} Y = (Y_{1}, Y_{2}, ..., Y_{\nu})' \in V' \\ |Y \sim N_{\nu}(\mu_{Y} = T_{Y}, \Sigma_{Y} = diag(\lambda_{1}, ..., \lambda_{\nu})) \end{cases}$$
(5)

$$Mp_{2} = P \begin{cases} Y = (Y_{1}, Y_{2}, ..., Y_{v}) \in V \\ |Y \sim N_{v}(\mu_{Y}, \Sigma_{Y} = diag(\lambda_{1}, ..., \lambda_{v})) \end{cases}$$
(6)

Y represents the vector of PC, Mp_1 is similar to MCp, Mp_2 is similar to MCpk. If $Mp_1 \ge 0.9973$, it indicates that this process has potential capability. If $Mp_2 \ge 0.9973$, it indicates that this process has actual capability.

3. The principal component contribution rate

The contribution rate of the principal component was determined by the variance ratio of the quality feature and the correlation coefficient. For convenience of explanation, the following two quality characteristics are taken as examples to describe and calculate the principle component contribution rate.

3.1 Theoretical analysis of the influencing factors of the principal component contribution rate

It is known from Figure 1 that when the correlation coefficients

of quality control features are same, the contribution rate of each principal component is only affected by the variance difference of each quality feature. Moreover, the higher the variance of quality characteristics, the higher the contribution rate of the first principal component. From the comparison of the three curves in Figure 1, it is found that with the increase of the variance ratio of the quality features, the correlation coefficient between quality characteristics has less influence on the contribution rate of the first principal component. When the standard deviation ratio of quality characteristics is more than 9 times, the contribution rate of the first principal component is independent from the correlation coefficient of the quality feature. What's more, it is only determined by the ratio of variance between quality characteristics. At this time, the contribution rate of the first principal component is approximately equal to the proportion of the maximum variance to the total variance.



Figure 1. Variance ratio of quality characteristic influence on principal component contribution degree

As is shown in Figure 2, when the variances of the control quality characteristics are same, the contribution rate of each principal component is affected by the correlation coefficient between the various quality features, which indicates that the greater the correlation coefficient between the quality features, the greater the contribution rate of the first principal component.

From the theoretical analysis above, during analyzing the covariance matrix based on quality features, the contribution rate of principal components is determined by the variance ratio and the correlation coefficient between quality features. However, during calculating the multivariate process capability index, multivariate process capability index is not only related to the variance and the correlation coefficient of quality characteristics, but also to the quality characteristic interval. Therefore, it is not advisable and reasonable to use the contribution rate of principal components to represent and evaluate the contribution rate of multivariate process capability

3.2 Application of the principal component contribution rate

The gearbox base used in this paper has two quality characteristics, the width X_1 (60mm) and the height X_2 (90mm). $S = \{(X_1, X_2), 57 \le X_1 \le 63, 89.4 \le X_2 \le 90.6\}$ is the specification range. We collected 100 sets of sample data of this part. The sample mean vector and the covariance matrix are obtained as:



Fig 2. Correlation coefficient of quality characteristic influence on principal component contribution degree

Through principal component analysis based on covariance matrix, the principal component contribution rates of each quality characteristic are obtained as:





As shown in Figure 3, $U=(u_1,u_2)$ is the eigenvector of S. The ellipse represents the area of 95% process capability, and the rectangular refers to the specification interval.

Obviously, the process capability of the first principal component is greater than that of the second principal component, which indicated that using the first principal component to calculate the multivariate process capability index will inevitably, overestimated the multivariate process capability. Because 91,7% of the first principal component contribution rate only explains 91,7% of the variation but not the multivariate process capability. It is ignoring the influence of the specification interval on the contribution rate of the multi-process capability, that leads to the over-estimation of the multi-process capability.

As illustrated by the above example, when calculating the multivariate process capability index, using the principal component contribution rate to evaluate the multivariate process capability directly is not precise. To solve this problem, we proposed a new method based on processing raw data.

4. MPCI with improved principal component contribution rate

Different from the traditional principal component analysis method, MPCI with improved principal component contribution rate needs to preprocess the quality features to standardize the specification intervals of each quality feature.

4.1 Data preprocessing

Interval standardization:

$$H_{i} = (X_{i} - M_{i}) / d_{i}, \quad i = 1, 2, ..., p$$
(7)

where X_i is the original quality feature, H_i is the processed quality feature, M_i is the center of the X_i specification interval, and d_i is half of the interval between the upper and lower specification boundaries of X_i :

$$M_i = \left(USL_{Xi} + LSL_{Xi}\right)/2 \quad d_i = \left(USL_{Xi} - LSL_{Xi}\right)/2$$

Through processing the original data, the specification intervals of all quality features are normalized to [-1, 1]. Thus, in the process of calculating the multivariate process capability index with the new data meter, the multivariate process capability index is only determined by the variance and correlation coefficient of the quality features.

In the process of principal component analysis of new data, because the principal component process with a residual contribution rate of less than 10% is very strong, these principal components can be ignored, and the total component is explained by a small part of the principal component whose cumulative contribution rate reaches 90% or more.

4.2 Multivariate process capability index calculation

For a part with p feature qualities $X_1, X_2, ..., X_p$, its feature quality vector is $X' = [X_1, X_2, ..., X_p]$, each of which has n sample points or observation data, and obeys $X \le N_p(\mu, \Sigma)$. O is the covariance matrix. Symbols used are defined as:

 $T'_X = (T_1, T_2, ..., T_P)$: X target value vector

$$LSL'_X = (LSL_1, LSL_2, ..., LSL_p)$$
: Lower specification limit of X

 $USL'_X = (USL_1, USL_2, ..., USL_P)$: Upper specification limit of X

 $S = \{x \mid LSL_X \leq x \leq USL_X\}$: X's super-rectangular specification interval

Quality characteristics H centralized treatment:

$$Z_{i} = H_{i} - \mu_{Zi}, \quad i = 1, 2, ..., p$$

$$So,$$

$$Z \le N_{p}(0, \Sigma_{z})$$
Where,
$$\Sigma_{z}(X_{i}, X_{z}) = \frac{Cov(X_{i}, X_{j})}{2}$$
(8)

$$\begin{array}{cccc}
\mathcal{L}_{z}(\Lambda_{i},\Lambda_{j}) & & \\
\mathcal{L}_{z}(\Lambda_{j},\Lambda_{j}) & & \\
\mathcal{L}_{z}(\Lambda_{j},\Lambda_{j$$

 $\lambda_1 \leq \lambda_2 \leq \dots \leq \lambda_p$: Characteristic value of Σ_z . u_1, u_2, \dots, u_p : Characteristic vector of Σ_z .

Principal component vector:

$$Y = U'Z$$
(10)
Where, $U = (u_1, u_2, ..., u_p), E(Y) = 0.$

Covariance of principal components:

$$COV(Y) = Diag(\lambda_1, \lambda_2, ..., \lambda_p)$$
⁽¹¹⁾

Covariance of principal components:

3

\$

$$T_{Zi} = (T_i - \mu_{Xi}) / d_i$$
Target value of quality characteristic Z_i: (12)

$$T_Y = U'T_Z \tag{13}$$

Y specification area:

$$R = \left\{ \left(y_1, y_2, \dots, y_{\nu} \right) \middle| LSL_z \le (U')^{-1} y \le USL_z, y' = \left(y_1, y_2, \dots, y_p \right), y_r = EY_r = 0 \right\} r = \nu + 1, \nu + 2, \dots, p$$
(14)

Probability that the principal component satisfies the specification area:

$$P_{I} = P \begin{cases} Y = (Y_{I}, Y_{2}, ..., Y_{v})' \leq R \\ |YnN_{v}(\mu_{Y} = T_{Y}, \Sigma_{Y} = diag(\lambda_{I}, ..., \lambda_{v})) \end{cases}$$

$$(15)$$

$$P_{2} = P \begin{cases} Y = (Y_{1}, Y_{2}, ..., Y_{\nu})' \in R \\ |Y \sim N_{\nu}(0, \Sigma_{Y} = diag(\lambda_{1}, ..., \lambda_{\nu})) \end{cases}$$
(16)

Get the new MPCI at this point:

$$Mcp = \frac{1}{3}\Phi^{-1}\left(\frac{P1+1}{2}\right)$$
(17)

$$Mcpk = \frac{1}{3}\Phi^{-1}\left(\frac{P2+1}{2}\right)$$
(18)

5. Application

Take the gearbox housing produced by a domestic manufacturer as an example. The proposed multi-process capability index (MPCI) is used for the capability analysis of the gearbox housing. In the manufacturing process, the specification intervals of the quality characteristics X_1 , X_2 , X_3 , and X_4 respectively are [-0.1, 0.1], [-0.035, 0.035], [-0.018, 0.004], and [-0.021, 0.004], and the target values are 0,0,-0.007 and -0.0085, respectively.

From the production process, 50 observation samples were randomly selected and according to the method of PAN et al.^[15], the Shapiro-Wilk statistic was calculated to check whether the sampled data follows the multivariate normal distribution. The Shapiro-Wilk statistic is 0.304078, which is greater than 0.05. Therefore, the sampled data is considered to follow a multivariate normal distribution.

Calculate the sample mean vector \overline{X} and the covariance matrix S:

$$\overline{X}' = \begin{bmatrix} 0.00841, 0.00281, -0.00743, -0.00821 \end{bmatrix}'$$

$$S = \begin{bmatrix} 0.000 \ 421 \ 33 \\ 0.000 \ 146 \ 98 \\ 0.000 \ 054 \ 25 \\ -0.000 \ 025 \ 18 \\ 0.000 \ 010 \ 52 \\ -0.000 \ 002 \ 71 \\ 0.000 \ 000 \ 89 \\ -0.000 \ 000 \ 14 \\ 0.000 \ 017 \ 28 \end{bmatrix}$$

5.1 MPCI with improved principal component contribution rate

Preprocess X according to equations (7) and (8):

$$H_i = (X_i - M_i) / d_i, \ Z_i = H_i - \mu_{Z_i}, \ i = 1, 2, ..., p$$

Calculate the Z-means vector \overline{Z} and the covariance matrix S₂ according to equation (9):

$$\overline{Z} = \begin{bmatrix} 0,0,0,0 \end{bmatrix}'$$

$$S_z = \begin{pmatrix} 0.042\ 133 \\ 0.041\ 994 & 0.068\ 473 \\ 0.049\ 318 & 0.065\ 403 & 0.086\ 942 \\ -0.002\ 168 & 0.002\ 034 & -0.001\ 018 & 0.110\ 592 \end{pmatrix}$$

Calculate each principal component according to equation (10):

$$\begin{split} \lambda_1 &= 0.1753 \\ PC_1 &= Y_1 = \mu_1 \leq Z = 0.4376Z_1 + 0.5883Z_2 + 0.6800Z_3 - 0.0069Z_4 \\ \lambda_2 &= 0.1107 \\ PC_2 &= Y_2 = \mu_2 \leq Z = 0.0168Z_1 - 0.0272Z_2 + 0.0027Z_3 - 0.9995Z_4 \\ \lambda_3 &= 0.0118 \\ PC_3 &= Y_3 = \mu_3 \leq Z = 0.3475Z_1 - 0.8076Z_2 + 0.4754Z_3 + 0.0292Z_4 \\ \lambda_4 &= 0.0104 \\ PC_4 &= Y_4 = \mu_4 \leq Z = 0.8292Z_1 + 0.0286Z_2 - 0.5582Z_3 + 0.0117Z_4 \end{split}$$

In addition, the principal components Y_1 , Y_2 , Y_3 , and Y_4 are independent of each other and obey the normal distribution, and their parameters are =(0, 0.1753), (0, 0.1107, (0,0.0118) , (0,0.0104) .The target vector of the principal component is (0.0422, -0.0283, 0.1361, -0.0375). The first two principal components explain the overall difference of 92.8%, so the first two principal components were used to analyze the multivariate process capability.

Since Z can be expressed as , according to the Z section the specification interval of the first two principal components can be calculated. Because only the first two principal components are considered, the expectation of the other two principal components is used when calculating the specification region. Its specification area is as follows:

$$R = \begin{cases} \left(Y_1, Y_2\right) & \frac{-0.1 + 0.00342}{0.1} \le 0.4376Y_1 + 0.01627Y_2 \le \frac{0.1 + 0.00342}{0.1} \\ \frac{-0.035 - 0.00298}{0.035} \le 0.5883Y_1 - 0.0272Y_2 \le \frac{0.035 - 0.00298}{0.035} \\ \frac{-0.018 + 0.00826}{0.011} \le 0.6800Y_1 + 0.0027Y_2 \le \frac{0.004 + 0.00826}{0.011} \\ \frac{-0.021 + 0.00889}{0.0125} \le -0.0069Y_1 - 0.9948Y_2 \le \frac{0.004 + 0.00889}{0.0125} \\ E(y_3) = 0, E(y_4) = 0 \end{cases}$$

As shown in Figure 4, the common area enclosed by 8 lines is the revised specification area:

$$R = \begin{cases} (Y_1, Y_2) \\ -0.969 \le -0.0069Y_1 - 0.9948Y_2 \le 1.031 \end{cases}$$

According to equation (11), (12) and (13), the target values of principal component covariance and Y are calculated as follow:

$$\Sigma_{Y} = \begin{pmatrix} 0.1753 & 0 \\ 0 & 0.1107 \end{pmatrix}$$
$$T_{Y} = (0.0422, -0.0283)$$

Calculate the probability that the principal component satisfies the specification region according to equations (15) and (16):

 $P_1 = 0.9965 P_2 = 0.9962$

MCP is calculated according to equation (17) and (18): Mcp,=0.9748

MCpk,=0.9651



Figure 4 Revised specification area

5.2 Traditional principal component method of MPCI

Calculate the principal component based on the covariance:

 $\lambda_1 = 0.0004839$

 $PC_{1} = Y_{1}' = -0.9287Z_{1}' - 0.3491Z_{2}' - 0.1250Z_{3}' + 0.0048Z_{4}'$ $\lambda_{2} = 0.0000299$ $PC_{2} = Y_{2}' = 0.3622Z_{1}' - 0.9068Z_{2}' - 0.1637Z_{3}' - 0.1404Z_{4}'$

 $\lambda_3 = 0.0000170$

 $PC_3 = Y'_3 = 0.0564Z'_1 - 0.1247Z'_2 - 0.0333Z'_3 + 0.9900Z'_4$ $\lambda_4 = 0.0000022$

 $PC_4 = Y_4' = 0.0562Z_1' + 0.2006Z_2' - 0.9780Z_3' - 0.0109Z_4'$

Principal components Y_{l^2} , Y_{2^2} , Y_{3^3} , and Y_4 are independent from each other and subject to normal distribution, and their parameters are μ , $\sigma^2 = (0,0.0004839)$, (0,0, 0.0000299 (0,0.0000170), and (0,0.0000022), respectively. The target vector of the principal component is (-0.0023, 0.0037, 0.0009, -0.0016). The first two principal components explain the overall difference of 96.4%, so the first two principal components were used to analyze the multivariate process capability.

$$R' = \begin{cases} \begin{pmatrix} Y_1', Y_2' \end{pmatrix} & -0.1 + 0.00342 \le -0.9287Y_1' + 0.3622Y_2' \le 0.1 + 0.00342 \\ -0.035 - 0.00298 \le -0.3491Y_1' - 0.9068Y_2' \le 0.035 - 0.00298 \\ -0.018 + 0.00826 \le -0.1250Y_1' - 0.1637Y_2' \le 0.004 + 0.00826 \\ -0.021 + 0.00889 \le 0.0048Y_1' - 0.1404Y_2' \le 0.004 + 0.00889 \\ & E(Y_3) = 0, E(Y_4) = 0 \end{cases}$$

As shown in figure 5, the common area surrounded by 8 lines is the revised specification area:

$$R' = \begin{cases} \left(Y_1', Y_2'\right) & -0.09658 \le -0.9287Y_1' + 0.3622Y_2' \le 0.10342 \\ -0.03798 \le -0.3491Y_1' - 0.9068Y_2' \le 0.03202 \\ -0.00974 \le -0.1250Y_1' - 0.1637Y_2' \end{cases}$$

The target values of principal component covariance and y': (0.0004839 0)

$$\Sigma_Y' = \begin{pmatrix} 0.0004039 & 0\\ 0 & 0.0000299 \end{pmatrix}$$

 $T_v = (-0.0023, 0.0037)^2$

The probability that the principal component satisfies the speification region:

$$P_1 = 0.99939$$

 $P_2' = 0.99936$

The MCPI:

Mcp₂=1.1428 MCpk₂=1.139



Figure 5 Revised specification area

The multi-process capability index should be less than the process capability index of any single quality feature. The potential process capability indices of the individual quality features X1, X2, X3, and X4 are calculated to be 1.62, 1.27, 1.13, and 1.002, so the multivariate process capability index should be less than 1.002.

In the example above, the potential MPCI with improved principal component contribution rate is 0.9748, and the potential MPCI of the traditional principal component method is 1.1428. The MPCI with improved principal component contribution rate is more realistic than the traditional principal component method. The analysis found that in the MPCI of the traditional principal component method, the variance of the quality feature X4 is smaller and the correlation is weaker than other quality features, so the contribution rate of the quality feature X4 is small and ignored (in the second section, it has been derived that the smaller the variance ratio and the weaker the contribution of the quality features, the smaller the contribution rate), resulting in MPCI being overestimated.

6. Conclusion

1) Through the theoretical analysis of the influence factors of the contribution rate of principal components, we found that the principal component contribution rate is determined only by the variance ratio and correlation coefficient between the quality features. Moreover, when one of them is constant, the contribution rate of the first principal component is positively linearly relative to another.

2) Based on the principal component method, we found

that the principal component contribution rate is not equivalent to the contribution rate of multivariate process capability practically. So, using the principal component contribution rate to directly evaluate the multivariate process capability is not reasonable or precise.

3) To solve the problem mentioned above, through normalizing the specification range of the processed quality features to [-1,1], eliminating the principal components with smaller variance to simplify the data. We proposed a method based on processing original data. Moreover, the rationality of the method is proved by an example.

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